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The effect of mindfulness-based interventions on stress and mental health outcomes among police officers: a systematic review and meta-analysis

Jennifer Gaskin ^{a*}, Wendy Wen Li ^{a*}, Liza van Eijk ^{a*} and Zoltan Sarnyai ^{b*}

^aCollege of Healthcare Sciences, James Cook University, Townsville, Australia; ^bCollege of Medicine and Dentistry, James Cook University, Townsville, Australia

ABSTRACT

The high demands of policing frequently result in stress, negatively impacting police officers' mental health. Mindfulness-Based Interventions (MBIs) have been increasingly recognised as effective for reducing stress in police. The aim of this systematic review and meta-analysis was to evaluate the effectiveness of MBIs on stress and mental health outcomes in police. Eleven studies were included, with a total of 967 participants. Both within- and between-group analyses demonstrated small to moderate effects of MBIs for reducing overall stress and improving mental health at post-intervention. These effects endured at follow-up for within-group comparisons. Between-group analyses showed MBIs had a small effect on overall mental health at follow-up in comparison to the control groups, however no significant effects on overall stress were detected at follow-up. MBIs may have a positive cumulative impact on stress and mental health, providing preliminary support for policing organisations to integrate MBIs into existing occupational health policy frameworks.

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Police; stress; mental health; mindfulness interventions; meta-analysis; systematic review


Introduction

Policing is an inherently stressful occupation, where cumulative exposure to acute and chronic stressors can have significant and adverse psychological impacts on those who serve in the profession. Beyond exposure to traumatic and critical incidents, studies indicate that operational and organisational factors, such as shift work, fatigue, public scrutiny, organisational culture, and administrative pressures contribute to stress experienced by officers (Carleton et al., 2020; Drew & Williamson, 2024; Purba & Demou, 2019). As a result, police officers experience persistently higher stress levels compared to other first responders and the general population (Kyron et al., 2022; Syed et al., 2020).

Stress occurs when the body's internal stability, referred to as homeostasis, is disrupted or challenged by an actual or perceived threat, triggering the stress response (Hendry et al., 2024;

CONTACT Jennifer Gaskin  jennifer.gaskin@my.jcu.edu.au  College of Healthcare Sciences, James Cook University, 1 James Cook Drive, Townsville QLD 4814, Australia

*My present affiliation is Margaret Roderick Centre for Mental Health Research, James Cook University, Australia.

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Lu et al., 2021). While acute stress promotes adaptation and is essential for successful coping, chronic stress, caused by prolonged exposure to stressors, is known to exacerbate pathogenic processes of disease (Creswell & Lindsay, 2014; Hendry et al., 2024), and is well associated with the onset of various mental health issues including depression, anxiety, post-traumatic stress disorder (PTSD), sleep disturbances, and increased alcohol use (Finlay et al., 2022; McEwen, 2017).

Empirical evidence among policing samples has demonstrated strong links between stress and psychological disorders, including anxiety, depression, and PTSD (Gershon et al., 2009; Gullon-Scott & Longstaff, 2024; Violanti et al., 2017; Zuiden et al., 2019). Increased alcohol consumption (Chopko et al., 2013; Ménard & Arter, 2013) and sleep disturbances (Garbarino et al., 2019; Violanti et al., 2017) also frequently present as comorbid conditions alongside high levels of stress among police. Stress from the daily work environment is more strongly associated with mental distress, sleep disturbances, and PTSD symptoms than exposure to traumatic and critical incidents (Drew & Williamson, 2024; Grupe, McGehee, et al., 2021). While stress can have a detrimental impact on the individual officer's psychological health and functioning, the consequences of chronic stress also permeate to policing organisations, resulting in increased absenteeism (Magnavita et al., 2013), use of excessive force (Christopher et al., 2024; Violanti et al., 2017), and premature resignation and retirement (Gullon-Scott & Longstaff, 2024), placing the organisation and community at risk. Providing officers with skills to cope with organisational and operational stressors is crucial for improving stress-related outcomes among police. Recognition of chronic stress and its detrimental impacts on the mental health of officers have prompted many policing organisations globally to prioritise and integrate interventions which are designed to reduce stress, and alleviate mental health symptoms in officers (Carleton et al., 2020; Moreno et al., 2024; Purba & Demou, 2019).

Mindfulness-based interventions (MBIs) have garnered significant scholarly attention for their potential to reduce stress and enhance mental health outcomes across various populations (Fendel et al., 2021; Gaskin et al. (in press); Li, McIntyre, et al. (2024); Ong et al. (2024); Qin et al. (2024); Sun et al. (2021); Xunlin et al. (2020)). Mindfulness involves directing purposeful attention to the present moment without judgment, fostering awareness and acceptance (Kabat-Zinn, 1982, 2013). Mindfulness can be cultivated through meditation, which can reduce negative affect and reactivity (Creswell & Lindsay, 2014; Nagy et al., 2020), in turn encouraging enhanced emotional regulation and promoting adaptive coping strategies towards stressful events (Crane et al., 2017; Kabat-Zinn, 2013, 2015; Li et al., 2023). First-generation MBIs are teacher-led programs involving structured training in both formal and informal meditation practices. These interventions have been extensively researched, providing the most robust evidence in the field of stress and mental health (Crane et al., 2017). In their meta-analytic review of randomised control trials (RCTs) in a population of active military personnel and veterans, Sun et al. (2021) found a small significant effect of MBIs for reducing military-related PTSD symptoms in comparison to controls. A meta-analysis examining the effects of MBIs on the mental health of nurses revealed medium-large effects on stress, and symptoms of depression, and anxiety, in comparison to control groups (Ong et al., 2024). The effects observed in doctors, however, were small and not significant (Ong et al., 2024), suggesting the effects of MBIs may vary in different populations, and by the nature of the stressors experienced. The effectiveness of MBIs may vary significantly when applied to policing populations compared to other groups due to the unique

demands and culture of police work. Policing involves continuous exposure to traumatic, stressful, and complex situations (Phythian et al., 2023), contributing to higher rates and severity of mental health challenges than seen in the general population (Kyron et al., 2022; Syed et al., 2020). Furthermore, the occupational culture of policing, characterised by stigma towards mental health, an emphasis on stoicism, and a strong professional identity tied to high-stakes decision-making (Drew & Martin, 2021), may influence the effectiveness of MBIs (Grupe, Smith, et al., 2021). Additionally, police training often reinforces suppression and desensitisation of trauma and stress, which may present specific challenges for the integration of mindfulness practices (Grupe, Smith, et al., 2021). These factors, distinct to the policing profession, suggest the potential for heterogeneity in outcomes when comparing the impact of MBIs on police to the effects in other populations. Therefore, it is crucial to advance the understanding of the effectiveness of MBIs specifically among policing samples.

The positive effects of MBIs on stress and mental health outcomes has prompted researchers to investigate the mechanisms underlying mindfulness. Self-awareness, emotion regulation, and attention have been proposed as key psychological mechanisms underscoring the effects of MBIs (Hölzel et al., 2011; Shapiro et al., 2018; Tang et al., 2015). From a neurobiological perspective, MBIs have been shown to alter brain structure and networks, mediating the positive effects on stress and mental health. Magnetic resonance imaging (MRI) studies have demonstrated increases in grey matter concentrations across various brain regions in participants post-MBI training compared to controls (Hölzel et al., 2011; Tang et al., 2015). This suggests that the effects of meditation may involve extensive brain networks. Specifically, focused attention meditation practices are believed to engage brain regions involved in attentional control, self-regulation (Ganesan et al., 2022; Hölzel et al., 2011; Tang et al., 2015), threat and stress responses, anxiety, and other mental health symptoms (Creswell & Lindsay, 2014; Eysenck et al., 2007; Williams et al., 2024). Among the policing population, MBIs have been shown to enhance attention (Hoeve et al., 2021; Navarrete et al., 2022), and non-reactivity as well (Christopher et al., 2016, 2018). Research also suggests that MBIs can be effective for reducing physiological stress reactivity across several interrelated physiological pathways associated with chronic stress (Gallistl et al., 2020; Gamaiunova et al., 2022; Rogerson et al., 2024), including modulating activity in the hypothalamic-pituitary-adrenal (HPA) axis, the body's primary stress response system responsible for detecting stressors, releasing hormones, and facilitating both the management and recovery from stressful situations. Given the well-established link between pathological functioning of stress symptoms and adverse psychological outcomes (Finlay et al., 2022; McEwen, 2017), understanding changes in physiological stress reactivity and biomarkers of stress is an important topic of enquiry among police. Findings from a recent meta-analytic review (Rogerson et al., 2024) indicated mindfulness and meditation-based interventions were effective for altering cortisol levels (Hedge's $g = .345$, $p < .001$) in other populations. Additionally, the findings of Grasmann et al. (2023) meta-analysis assessing the effects of MBIs on biomarkers of stress in psychiatric, stressed, and healthy adults demonstrated MBIs ameliorated biological markers of stress at post-intervention, with the magnitude of effects increasing at follow-up. Biological measures of stress have been incorporated into studies to provide complimentary data for the impacts of MBIs on stress in police, with studies reporting increased short- and long-term heart rate variability (HRV; Krick

& Felfe, 2020), and significantly lower cortisol awakening response (CAR) relative to waitlist controls at three month follow-up (Grupe, Stoller, et al., 2021). The results, however, have been inconsistent, with some research indicating no significant effects of MBIs on changes in CAR (Christopher et al., 2016, 2018), hair cortisol concentrations (HCC; Grupe, Stoller, et al., 2021), or heart rate (Krick & Felfe, 2020).

A range of first-generation MBIs (Crane et al., 2017) have begun to be utilised among samples of police to examine the effects of MBIs on stress and mental health outcomes. The clinically standardised Mindfulness-Based Stress Reduction (MBSR) program (Kabat-Zinn, 1982) is a non-trauma focused group treatment which is delivered over eight weeks. Throughout the program, participants learn to attend to the present moment through a combination of didactic and experiential methods, somatic exercises, and meditation practices (Kabat-Zinn, 1982, 2013). Mindfulness-Based Resilience Training (MBRT; Christopher et al., 2016) places emphasis on normalising the effect of stress and trauma inherent to police work and integrates the context of policing and policing culture into a group-based mindfulness program. Mindfulness-based Cognitive Therapy (MBCT; Segal et al., 2002) a program which focuses on re-evaluating maladaptive thought patterns to reduce symptoms of anxiety, and depression, and other variants of MBIs have also been utilised within policing samples. Studies within policing populations have shown significant reductions in perceived, operational, and organisational stress, anxiety, depression, PTSD, and sleep disturbances following MBIs (Christopher et al., 2018; Grupe, McGehee, et al., 2021; Khatib et al., 2022; Márquez et al., 2021; Navarrete et al., 2022; Trombka et al., 2021). However, the findings across studies involving police are mixed. Christopher et al. (2018) found no significant effects of MBIs for reducing operational stress, anxiety or depression at post-intervention or follow-up relative to controls. Other studies have reported no significant reductions in organisational, and perceived stress, sleep disturbances, depression (Grupe, Stoller, et al., 2021), and alcohol use (Christopher et al., 2018) at post-intervention, as well as organisational stress, and sleep disturbances at follow-up (Christopher et al., 2018; Grupe, Stoller, et al., 2021).

While a growing number of studies have explored the effects of MBIs on stress and mental health outcomes in police, this research is still at an early stage. Moreover, the findings of the studies lack consistency, exemplifying the need to systematically examine this emergent field of research. Variability in results can be attributed to several factors, including the specific MBI protocol employed, such as duration, delivery format, or components, which can may influence outcomes (Melis et al., 2022; Moyes et al., 2022). The country of study also plays a crucial role, as cultural, societal, and systemic differences – such as attitudes toward mental health, accessibility of support services, and police organisational norms – may shape the impact of MBIs on policing populations (Grupe, Smith, et al., 2021; Kabat-Zinn, 2019). Additionally, the measures used in studies and the timing of the study can impact the results (APA, 2020). These considerations underscore the importance of context in understanding the effectiveness of MBIs among police. On this basis, this study will explore whether the effect differs across MBI protocols, country, and scales used to assess outcomes to identify potential moderators of the effect of MBIs on police stress and mental health.

There are limited reviews which provide a clear and comprehensive synthesis of the available literature. A search in Cochrane, PROSPERO and eight databases (CINAHL, Medline, PsychINFO, PubMed, EMCARE, Web of Science, SCOPUS, and Criminal

Justice Database) located one meta-analysis (Withrow et al., 2024) and one systematic review (Vadvilavičius et al., 2023) evaluating the effectiveness of MBIs among police officers. Vadvilavičius et al. (2023) systematic review suggests MBIs are useful for reducing perceived stress, and improving mental health outcomes among police, however, no meta-analysis was undertaken to quantify the effectiveness of the programs. The findings of Withrow et al. (2024) meta-analysis indicated moderate post-intervention effects of MBIs on depression ($SMD = .56$, $CI\ 95\% [.94, .19]$, $p = .003$), and small post-intervention effects on anxiety ($SMD = .32$, $95\% CI [.64, .00]$, $p = .05$). The meta-analysis reported effects sizes for the outcomes of depression and anxiety, however the data utilised in the analyses was measured in the original study utilising the Negative Affect Scale, and the Psychological Strain Scale. These scales do not exclusively measure the outcomes of depression and anxiety, therefore the effects observed in the meta-analysis may not accurately reflect the effect of MBIs on depression and anxiety specifically. Moreover, the study did not quantify the effects of MBIs on officer perceived stress, nor did it extend its evaluation to the effects of MBIs on other highly prevalent mental health conditions among police, including PTSD, and alcohol use (Chopko et al., 2013; Kyron et al., 2022). Neither Vadvilavičius et al. (2023) or Withrow et al. (2024) study quantified the effect of MBIs on biological measures of stress across studies involving police. Lastly, the two existing studies did not quantify the medium- and long-term effectiveness of MBIs on stress and mental health outcomes – an essential factor in assessing the sustained efficacy and feasibility of MBIs among police. A comprehensive assessment of the effects of MBIs for attenuating stress and improving mental health in police is needed to advance and guide future research, practice and policy.

The aim of this systematic review and meta-analysis is to address these literary gaps, and to evaluate whether MBIs are effective for reducing stress (police-specific organisational and operational stress, and general perceived stress) and improving mental health outcomes (anxiety, depression, PTSD, sleep disturbances, and alcohol use) in police officers. In addition to examining the effects of MBIs on individual stress and mental health outcomes in police officers, this study also considers the complex and interconnected nature of these issues by evaluating the combined effects of MBIs on stress and mental health overall. By pooling outcomes, this approach reflects the real-world complexity of stress and mental health challenges experienced by police.

As such, the current study will answer the following research questions (RQ):

RQ1: Are mindfulness-based interventions effective for reducing stress in police officers?

RQ2: Are mindfulness-based interventions effective for reducing mental health symptoms in police officers?

Method

This systematic review and meta-analysis was registered in PROSPERO under the registration number CRD42024538639 and adhered to the PRISMA guidelines

(Preferred Reporting Items for Systematic Reviews and Meta-Analyses).

Inclusion and exclusion criteria

Inclusion criteria for the current study were quantitative studies investigating the effectiveness of mindfulness-based interventions on stress and/or mental health outcomes of depression, anxiety, PTSD, alcohol use, and sleep disturbance among sworn police officers. In regard to outcomes, restrictions were imposed to ensure the accuracy and consistency of this meta-analysis findings. Firstly, the measurement tools used in the included studies needed to specifically measure the intended outcome (Deeks et al., 2024; Murad et al., 2019). For instance, to be eligible under the outcome of depression, the measurement scale needed to be explicitly designed to assess depression, such as the Beck Depression Inventory (BDI) or the Patient Health Questionnaire-9 (PHQ-9). Secondly, the included studies were restricted to those that measured outcomes at a continuous or interval level, providing detailed and precise measurements. This approach allowed us to standardise and pool the data effectively, thereby facilitating comparability across studies (Borenstein et al., 2019; Deeks et al., 2024). Studies were also eligible for inclusion if they were published in peer-reviewed English language journals and comprise random controlled trials (RCTs), non-random controlled trials (NRCT), and single arm studies. First-generation MBIs represent the most robust evidence in the field, therefore this systematic review and meta-analysis considered studies which implemented teacher-led first-generation MBIs supporting the sustained and systematic training of mindfulness meditation practices.

The exclusion criteria were studies: 1) involving retired police or unsworn civilian employees where separate data for sworn police were not available; 2) studies which did not employ a teacher-led first wave MBI (e.g., self-paced online program/app, mindfulness-informed treatment, unstructured mindfulness programs or practices; 3) without full text; 4) case studies, thesis submissions, systematic and meta-analytic reviews, book chapters, conference presentations, and non-empirical articles (e.g., editorials, opinion pieces, reviews).

Search strategy

Eight online databases (Medline, EMCARE, Cochrane, CINAHL, PsychINFO, PubMed, SCOPUS, and Web of Science) were independently searched by JG and WL between 5th and 8th of April 2024 for articles published from the inception of each database to April 2024. The search was repeated by JG on 15 November 2024. No new articles which met inclusion criteria were identified. In addition, a search of Criminal Justice Database (ProQuest) was conducted by JG on 23 March 2025 for articles published from database inception to March 2025. No new articles meeting inclusion criteria were located in the search. Medical Subject Heading (MeSH), keywords and suggested subject terms were utilised during the search. The full search strategy utilising Cochrane's PICO search tool (Higgins et al., 2022) is presented in Tables S1-S4 of the Supplementary Materials.

Study selection

Following the removal of duplicates, the remaining articles underwent review and double screening by the first two authors. The first stage of study selection involved title and abstract screening against exclusion and inclusion criteria. All title and abstracts were independently reviewed by JG and WL and were coded using 'yes', 'no' and 'maybe'. Studies coded unanimously as 'yes', or 'maybe' were included for full text review, and studies coded unanimously as 'no' were excluded. Articles which received a nonunanimous coding were reviewed and discussed until an agreement was achieved for full-text assessment (Fisher et al., 2023; Leow et al., 2024). Each eligible full-text article was then reviewed against inclusion and exclusion criteria, and coded 'yes', 'no' and 'maybe' by both JG and WL independently. Studies unanimously coded as 'yes' were included for further assessment, and studies coded unanimously as 'no' were excluded. Articles which received a non-unanimous coding or a 'maybe' were discussed until an agreement was achieved for either inclusion or exclusion (Fisher et al., 2023; Leow et al., 2024).

The second stage of selection involved the methodological appraisal of the full text of each included article utilising the Mixed Methods Appraisal Tool (MMAT) Version 18 (Hong et al., 2019). Each full text article was appraised and coded by both JG and WL independently. The MMAT includes a set of criteria and screening questions to assess the quality of empirical studies, providing an overall quality score. Two initial questions determine if the study is empirical and suitable for appraisal using the MMAT. Two reviewers then assess the study's quality against five specific criteria. Each criterion is coded as 'Yes' (criterion met), 'No' (criterion not met), or 'Can't tell' (insufficient information). Cohen's Kappa (k) was used to identify the inter-rater agreement indices, where values of k represented poor agreement ($k < .20$), fair agreement ($k = .40$), moderate agreement ($k = .60$), substantial agreement ($k > .80$), and perfect agreement (Cohen, 1968). Studies rated as $> .40$ were automatically included (Landis & Koch, 1977; M. Li et al., 2023; McHugh, 2012). As suggested by the manual of MMAT (Hong et al., 2018), studies with low methodological quality were not automatically removed, they instead were discussed and re-evaluated. Two studies (Khatib et al., 2022; Krick & Felfe, 2020) received inter-rater agreement value $k < .40$. After discussion between JG and WL, it was agreed that both studies would be included for data extraction. Table S5 in the Supplementary Materials displays the results of the MMAT and inter-rater agreement indices.

Data extraction

Data was extracted from eligible studies utilising an Excel spreadsheet to extract first author, citation, publishing year, country of the study, sample size, demographics, study design, MBI protocol, control protocol, measures and findings of subjective stress, biological stress, and mental health variables. Study data was extracted independently by JG and WL and then reviewed by each other to assess the accuracy of data extraction. Any inaccuracies were discussed and rectified between JG and WL.

The extracted data was assessed and coded independently by JG and WL to evaluate if the interpretation of the findings presented in the study was supported by the results of the statistical analysis. Coding of 'unsupported' (= 0), 'credible' (= .5), or 'unequivocal'

(= 1) was used. An agreement index for each article was then calculated $((N_{\text{unequivocal}} + N_{\text{credible}})/N_{\text{reviewers}})$. Articles were eligible for data synthesis if the agreement index was higher than 0.80 (Astridge et al., 2023; Freedman et al., 2024; Li, Nannestad, et al., 2024). All studies achieved agreement indices higher than 0.80 therefore all studies were included in the meta-analysis.

Data synthesis

Meta-analysis and narrative synthesis were utilised to synthesise the data. Meta-analysis was performed utilising Comprehensive Meta-Analysis v4 (CMA v4; Borenstein et al., 2019).

Individual effect sizes were calculated using Hedge's g in CMA. Hedges' g was utilised for this study due to its ability to correct for small sample sizes and its robustness in handling heterogeneity in both sample characteristics and measurement scales, ensuring accurate and standardised effect size estimates (Cohen, 1977; Hedges, 2019). To calculate the overall effect size across studies, a random effects model was used (Borenstein et al., 2021; Hedges, 2019). Included studies reported data on several outcome variables based on the same participants, therefore, pooled effect sizes were calculated using a multiple outcomes model, accounting for correlations between different outcomes within the same study in the modelling approach (Borenstein, 2019; Borenstein et al., 2021). The correlations between different outcomes within the same study were accounted for by assuming a correlation of $r = 0.5$ (Wampold et al., 1997; White et al., 2019; Wood et al., 2023).

Within-group comparisons included data of the MBI intervention group of included studies. Within-group mean differences were calculated by deducting the baseline means from post-intervention and follow-up means. Hedges' g was calculated using CMA, which automatically computed the effect sizes based on the input means, standard deviations, and sample sizes.

While within-group comparisons have limitations, due to the infancy of research on MBIs for police officers, within-group comparisons were conducted and included in this study. Conducting RCTs in policing is often challenging due to high costs and logistical complexities which may not be feasible in all settings and countries. Within-group comparisons provide valuable evidence for future studies using single-arm designs, crucial for extending the evidence base for MBIs' feasibility and effectiveness across different countries and police subpopulations. These comparisons allow for examining intervention efficacy in real-world contexts where control groups may not be ethical or feasible. Given the infancy of MBI research in policing, within-group comparisons offer valuable preliminary evidence, helping to build a foundation for future, more extensive RCTs. This approach ensures that even in the absence of RCTs, valuable insights can be gained, by allowing the assessment of changes over time within the same group (Deeks et al., 2024; Evans, 2010).

For between-group analyses RCTs and NCRTs were included to provide a comprehensive assessment of the available evidence, where an MBI intervention group and an inactive control group (waitlist [WL] or treatment as usual) were employed. Incorporating both RCTs and NRCTs enhances the generalisability findings and allowed sensitivity analyses to be performed to assess the robustness of the results (Borenstein et al., 2021; Mei et al., 2024). The group differences were calculated by deducting the mean

scores of the outcome variable of the control group at post-intervention and follow up, from the mean scores of the MBI intervention group. Individual effect sizes were first calculated using Hedge's g in CMA, and then a random effects model was used to pool effect size across studies. Two studies (Grupe, McGehee, et al., 2021, 2021) reported multiple effect sizes for sleep outcomes (disturbances and quality), which were not independent, therefore a two-level meta-analysis was utilised (Li, Nannestad, et al., 2024). First, employing a fixed effect model, a synthetic effect size was computed for the studies (Tables S6-S9). Next, the synthetic effect was input into the main meta-analysis to compute the overall effect size across studies using the random effects model (Borenstein et al., 2021; Hedges, 2019).

To report pooled effect sizes, Hedges' g with 95% confidence interval (95% CI) was utilised, with values of 0.20, 0.50, and 0.80 representing small, medium, and large effect sizes, respectively. Between study heterogeneity was assessed with I^2 statistic. The magnitude of heterogeneity was categorised as low, medium, and high represented by $I^2 = 25, 50, \text{ and } 75$ or greater (Borenstein, 2019). If medium or high heterogeneity was detected, meta-regression was administered to determine which factors contributed to the heterogeneity. Publication bias was evaluated using multiple methods, including Egger's regression test (Egger et al., 1997), which indicated bias when significant p -values were observed (Borenstein, 2019). Additionally, Rosenthal's failsafe N (Rosenthal, 1979), Duval and Tweedie's trim-and-fill method (Duval & Tweedie, 2000), and the visual inspection of funnel plots were employed to further assess publication bias and to identify missing studies.

Assessing risk of bias in included studies

An assessment of the risk of bias (ROB) was performed independently by JG and WL using the MMAT, where 'yes' was coded as 1, 'can't tell' coded as 0.5, and 'no' coded as 0 (Emary et al., 2022) for each of the seven MMAT assessment items. The maximum score for each article in this study was 14 (two raters; JG and WL). The ROB index was calculated as a percentage of the maximum score, where $> 0.80 = \text{low risk}$, $0.40\text{--}0.79 = \text{medium risk}$, and $< 0.40 = \text{high risk}$ (Wigg et al., 2024). For all included studies, the overall risk of bias was rated as low.

Results

As shown in the PRISMA diagram (Figure 1), 11 studies were included in the systematic review, and ten studies in the meta-analysis. Four articles (Imhoff-Smith & Grupe, 2023; Kaplan et al., 2020; Rehder et al., 2021; Ribeiro et al., 2020) were secondary analyses of two primary studies (Christopher et al., 2018; Grupe, Stoller, et al., 2021) utilising the same participants and outcome variables as the primary studies, therefore were regarded as the same study. Study data were requested from the authors of three studies (Christopher et al., 2024; Grupe, Stoller, et al., 2021; Khatib et al., 2022) for inclusion in the meta-analysis.

Summary of included studies

Of the included studies, six were conducted in the U.S.A. ($n = 6$), Spain ($n = 2$), Netherlands ($n = 1$), Germany ($n = 1$), and Brazil ($n = 1$). The combined sample size

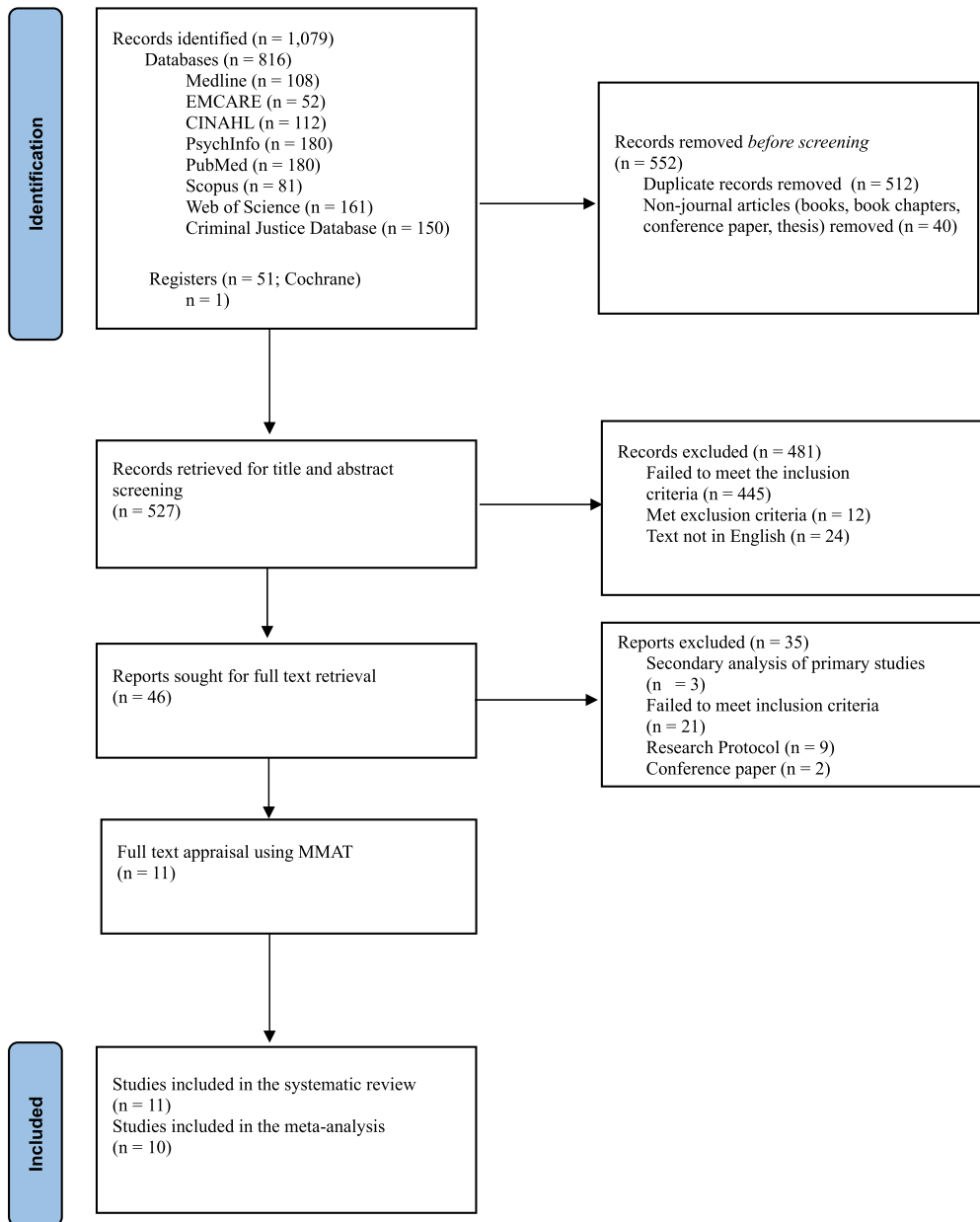


Figure 1. PRISMA diagram.

included 967 participants, with sample sizes of individual studies ranging from 20 to 267. Male participants made up 61% ($n = 609$), and sample age ranged from 18 to 65 years.

Of the included studies, one study employed clinically standardised MBSR (Khatib et al., 2022), three implemented MBRT (Christopher et al., 2016, 2018; Khatib et al., 2022), two utilised a low-dose MBSR protocol (Krick & Felfe, 2020; Márquez et al., 2021), and the remaining delivered other MBI variants (Christopher et al., 2024; Grupe, McGehee, et al., 2021; Grupe, Stoller, et al.,

2021; Hoeve et al., 2021; Khatib et al., 2022; Navarrete et al., 2022; Trombka et al., 2021). Two studies reported the MBI was delivered online (Christopher et al., 2024; Khatib et al., 2022). A summary of the characteristics of the included studies is presented in Table 1.

Analysis of RQ1

RQ1: Are mindfulness-based interventions effective for reducing stress in police officers?

Between-group effects comparing MBI and control groups at post-intervention

The forest plot of the analysis is displayed in Figure 2. A small significant effect of MBIs on overall stress favouring treatment groups compared to the control groups was found (Hedge's $g = -.198$, 95% CI $[-.353, -.043]$, $p = .012$). Small to moderate effects of MBIs for individual stress outcomes of operational stress (Hedge's $g = -.104$, 95% CI $[-.398, .190]$, $p = .488$), organisational stress (Hedge's $g = -.268$, 95% CI $[-.688, .152]$, $p = .210$), perceived stress (Hedge's $g = .170$, 95% CI $[-.448, -.108]$, $p = .231$), and cortisol awakening response Hedge's $g = -.289$, 95% CI $[-.584, .007]$, $p = .055$) were found, however significance was not reached.

The I^2 statistics indicated an absence of heterogeneity for cortisol awakening response ($I^2 = .000$, $p = .695$), operational stress ($I^2 = .000$, $p = .945$), and perceived stress ($I^2 = .000$, $p = .529$), and low but non-significant heterogeneity for organisational stress ($I^2 = 46.641$, $p = .171$). Meta-regression was therefore not performed. Publication bias was not detected, indicated by non-significance of Egger's regression test (intercept = $-.266$, $t = .205$, $df = 7$, $p = .843$). Inspection of the funnel plots, and Duval and Tweedie's trim-and-fill analysis revealed one missing study with the observed effect altered from Hedge's $g = -.198$, 95% CI $[-.353, -.043]$ to Hedge's $g = -.217$, 95% CI $[-.358, -.076]$, however the magnitude and directions of the effect remained unchanged. The funnel plot is presented in Figure S1 of the Supplementary Materials.

Between-group effects comparing MBI and control groups at follow-up

As displayed in Figure 3, the overall summary effect of MBIs on stress small (Hedge's $g = -.094$, 95% CI $[-.282, .094]$, $p = .327$), but did not reach significance. The pooled effect sizes indicated small effects of MBIs on operational stress (Hedge's $g = -.033$, 95% CI $[-.327, .261]$, $p = .825$), organisational stress (Hedge's $g = -.079$, 95% CI $[-.373, .215]$, $p = .597$), and perceived stress (Hedge's $g = -.265$, 95% CI $[-.706, .176]$, $p = .239$), however the effects were not significant in comparison to the control groups at follow-up.

An absence of heterogeneity was indicated by a zero I^2 statistic for both operational stress ($p = .930$), and organisational stress ($p = .722$), and low but non-significant heterogeneity was found for perceived stress ($I^2 = 42.060$, $p = .189$), therefore meta-regression was not performed. Egger's regression test did not detect publication bias (intercept = 1.288 , $t = .622$, $df = 4$, $p = .568$). Duval and Tweedie's trim-and-fill analysis and inspection of the imputed funnel plot identified one missing study, which increased the effect from Hedge's $g = -.094$, 95% CI $[-.282, .094]$ to Hedge's $g = -.148$, 95% CI $[-.314, .018]$,



Table 1. Summary of the characteristics of included studies investigating mindfulness-based interventions among police officers.

Author	Country	Study Design	MBI Details	Participant Demographics	Measures
Christopher et al. (2018)	USA	RCT A1 = MBRT A2 = WL	MBRT 8 weeks 2 hr weekly sessions 2 hr retreat week 7 Home practice activities Face-to-face delivery	N = 61 (Intervention group: n = 31; control group: n = 30) Age: Intervention M = 44.73 (SD = 6.63); Control M = 43.22 (SD = 5.43) Sex: 7 female; 54 male Years in job: Intervention M = 18.50 (SD = 6.98); control M = 17.97 (SD = 6.69)	PROMIS-SD PROMIS-AU PROMIS-A PROMIS-D PSQ CAR
Christopher et al. (2016)	USA	Single Arm A1 = MBRT	MBRT 8 weeks 2 hr weekly sessions 6 hr retreat week 7 Home practice activities Face-to-face delivery	N = 62 Age: M = 43.76 (SD = 7.22) Sex: 22 female; 40 male Years in job: M = 13.72 (SD = 5.96)	PROMIS-SD PSS PSQ CAR
Christopher et al. (2024)	USA	RCT A1 = MBRT A2 = SME A3 = WL	MBRT/Modified MBRP 8 weeks 6 hr intensive week 1 1 hr weekly sessions 4 hr retreat week 7 Home practice activities Online delivery	N = 73 (Intervention group 1: n = 33; intervention group 2: n = 24; control group 2: n = 16) Age: Intervention 1: M = 39.57 (SD = 10.28); intervention 2: M = 38.08 (SD = 6.80); control M = 37.06 (SD = 10.14) Sex: 22 female; 51 male Years in job: Intervention 1: M = 11.82 (SD = 9.51); intervention 2: M = 10.48 (SD = 5.62); control M = 9.03 (SD = 9.05)	PROMIS-SD-SF PROMIS-AU-SF PROMIS-D-SF PROMIS-A-SF PCL-5 PSS
Grupe, McGehee, et al. (2021)	USA	Single Arm A1 = MBI	Modified MBSR/MBRT 8 weeks 2 hr weekly sessions 4 hr retreat week 7 Home practice activities Face-to-face delivery	N = 30 Age: M = 38.4 (SD = 7.70) Sex: 16 female; 14 male Years in job: M = 10.80 (SD = 7.50)	PSQ PSQI PCL-C PROMIS-D PROMIS-A PROMIS-SD

(Continued)

Table 1. (Continued).

Author	Country	Study Design	MBI Details	Participant Demographics	Measures
Grupe, Stoller, et al. (2021)	USA	RCT A1 = MBI A2 = WL	Modified MBSR/MBRT 8 weeks 2 hr weekly sessions 4 hr retreat week 7 Home practice activities Face-to-face delivery	N = 114 (Intervention group: n = 57; control group: n = 57) Age: Intervention; M = 40.20 (SD = 7.40); control M = 39.80 (SD = 9.30) Sex: 47 female; 67 male	PSS PSS PTSD-Checklist PROMIS-A PROMIS-D PROMIS-SD PSQI AUDIT CAR HCC
Hoeve et al. (2021)	Netherlands	Single Arm A1 = MBI	MBI based on Williams & Cullen, 2013 6 weeks 2 hr weekly sessions No retreat Home practice activities Face-to-face delivery	Years in job: Intervention M = 13.8 (SD = 7.90); control M = 14.30 (SD = 8.30) N = 82 Age: M = 49 (SD = 8.40) Sex: 46 female; 36 male Years in job: Unreported	DASS-21 (stress subscale only) PSQ-Op SCL-90
Khatib et al. (2022)	USA	RCT A1 = MBSR A2 = MBRT A3 = mPEAK	MBSR 8 weeks 2.5 hr weekly sessions 7 hr retreat Home practice activities Online delivery MBRT 8 weeks 2.5 weekly sessions 7 hr retreat Home practice activities Online delivery mPEAK 8 weeks 2.5 hr weekly sessions 7 hr retreat Home practice activities Online delivery	N = 50 (Intervention group 1: n = 14; intervention group 2: n = 17; intervention 3: n = 19) Age: M = 44 (SD unreported) Sex: 16 female; 34 male Years in job: Unreported	PSS BDI-II SAI

(Continued)



Table 1. (Continued).

Author	Country	Study Design	MBI Details	Participant Demographics	Measures
Krick and Felfe (2020)	Germany	RCT A1 = MBI A2 = TAU	Modified MBSR 6 weeks 2 hr weekly sessions Retreat; not reported Home practice activities Face-to-face delivery	N = 267 (Intervention group: n = 126; control group: n = 141) Age: M = 25.96 (SD = 5.57) Sex: 57 female; 210 male Years in job: Unreported	HR HRV
Márquez et al. (2021)	Spain	Single arm A1 = MBI	Modified MBSR 7 weeks 2 hr weekly sessions x 6 4 hr retreat Home practice activities Face-to-face delivery	N = 20 Age: M = 45.63 (SD = 10.17) Sex: 11 female; 9 male Years in job: Unreported	PSS
Navarrete et al. (2022)	Spain	NRCT A1 = MBI A2 = WL	MPPW 8 weeks 2 hr weekly sessions Retreat; not reported Home practice activities Face-to-face delivery	N = 38 (Intervention group: n = 20; control group: n = 18) Age: M = 39.97 (SD = 6.41) Sex: 15 female; 23 male Years in job: Unreported	DASS-21 PROMIS-SD
Trombka et al. (2021)	Brazil	RCT A1 = MBI A2 = WL	MBHP 8 weeks Session length; not reported Retreat; not reported Home practice activities Face-to-face delivery	N = 170 (Intervention group: n = 88; control group: n = 82) Age: M = 42.26 (SD = 7.71) Sex: 127 female; 43 male Years in job: Unreported	HADS

MBRT = Mindfulness-Based Resilience Training; PROMIS = Patient-Reported Outcomes Measurement Information System; PROMIS-SD = PROMIS Sleep Disturbance; PROMIS-AU = PROMIS Alcohol Use; PROMIS-A = PROMIS Anxiety; PROMIS-D = PROMIS Depression; PSQ = Police Stress Questionnaire; CAR = Cortisol Awakening Response; PSS = Perceived Stress Scale; SME = Stress Management Education; MBRP = Mindfulness-Based Relapse Prevention; SF = Short Form; PCL-5 = PTSD Checklist for DSM-5; MBSR = Mindfulness-Based Stress Reduction; PSQI = Pittsburgh Sleep Quality Index; PCL-C = PTSD Checklist-Civilian Version; AUDIT = Alcohol Use Disorder Identification Tool; HCC = Hair Cortisol Concentration; MBI = Mindfulness-based intervention; DASS-21 = Depression Anxiety Stress Scale; SCL-90 = Symptom Checklist-90; mPEAK = Mindful Performance Enhancement, Awareness, and Knowledge; BDI-II = Beck Depression Inventory-II; SAI = State Anxiety Inventory; HR = Heart rate; HRV = Heart rate variability; MBHP = Mindfulness-Based Health Promotion; HADS = Hospital Anxiety and Depression Scale.

^aSix month follow-up.

^bSix week follow-up.

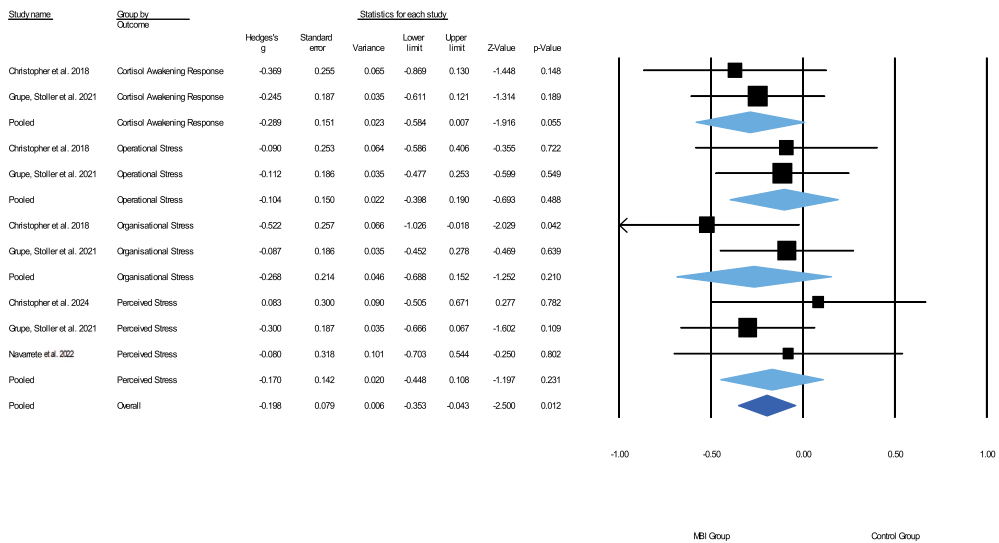


Figure 2. Forest plot of between-group analysis comparing MBI and Control groups at post-intervention on stress.

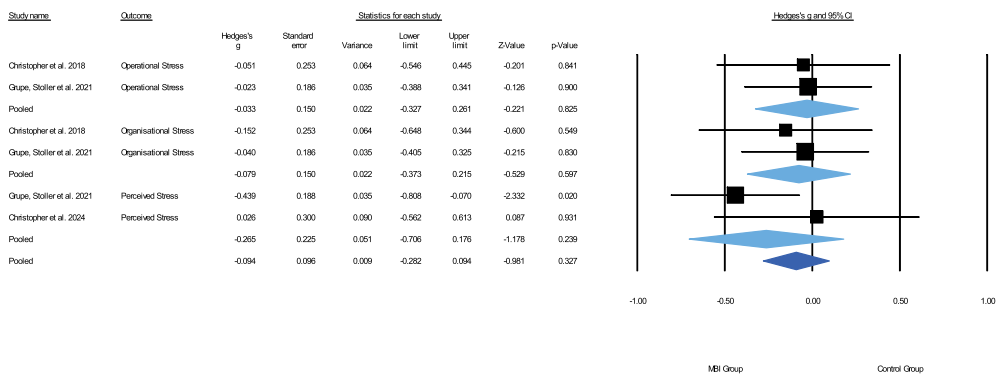


Figure 3. Forest plot of between-group analysis comparing MBI and Control groups at follow-up on stress.

however the direction and magnitude of the effect remained unchanged. Figure S2 of the Supplementary Materials displays the funnel plot.

Sensitivity analysis

To determine the degree of robustness of the findings, a sensitivity analysis was performed. Removing one study comprising the NRCT study design (Navarrete et al., 2022) did not change the direction or significance of the effect at post-intervention. The results of the sensitivity analyses are shown in Table S10 of the Supplementary Materials.

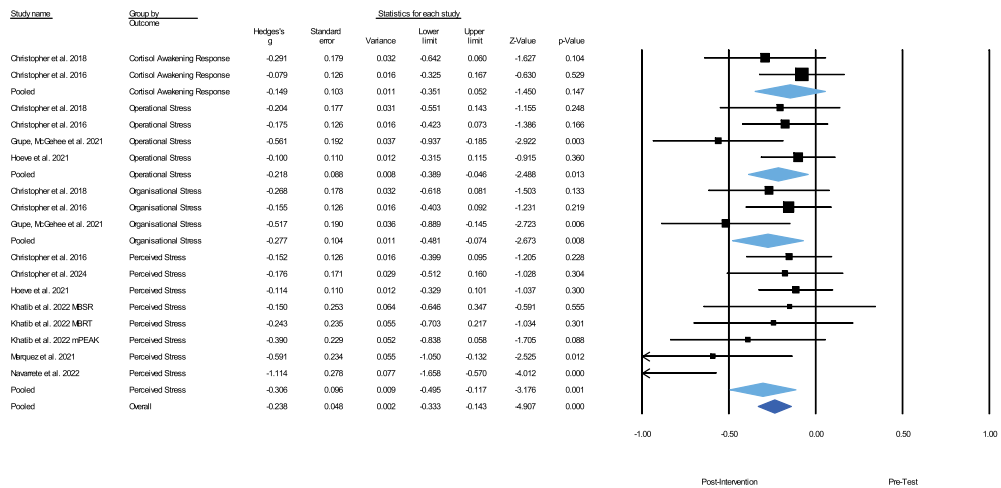


Figure 4. Forest plot of within-group analysis at post-intervention on stress.

Within-group effects comparing post-intervention against pre-intervention

The overall summary effect at post-intervention on stress was small (Hedge's $g = -0.238$, 95% CI $[-0.333, -0.143]$, $p = .000$). The forest plot of the analysis showing the pooled effects for individual outcomes and studies is displayed in Figure 4.

As indicated by the I^2 statistics an absence of heterogeneity was found for CAR ($p = 0.332$), and low but non-significant heterogeneity for operational stress ($I^2 = 31.793$, $p = 0.222$), and organisational stress ($I^2 = 20.588$, $p = 0.284$). Moderate heterogeneity was evident for perceived stress ($I^2 = 52.240$, $p = .041$). The moderators of publication year ($Q = 2.92$, $df = 1$, $p = .087$), MBI protocol ($Q = 3.10$, $df = 3$, $p = .377$), and measure ($Q = 1.07$, $df = 3$, $p = .785$) were not responsible for the heterogeneity. The moderators of sample size ($Q = 20.91$, $df = 8$, $p = .007$), country ($Q = 12.93$, $df = 2$, $p = .002$), and research design ($Q = 10.90$, $df = 2$, $p = .004$) were accountable for the significant heterogeneity. The results of the meta-regression are displayed in Table S11 of the Supplementary Materials. Egger's regression test evidenced publication bias (intercept = -3.143 , $t = 4.440$, $df = 15$, $p < .001$). Duval and Tweedie's trim-and-fill analysis and inspection of the adjusted funnel plot identified seven missing studies. Imputation of these studies reduced the observed effect from Hedge's $g = -0.238$, 95% CI $[-0.333, -0.143]$ to Hedge's $g = -0.144$, 95% CI $[-0.253, -0.035]$. Inspection of the fail-safe N values (Rosenthal, 1979) indicated 177 additional studies with null results would be required to nullify the effect. The funnel plot is displayed in Figure S3 of the Supplementary Materials.

Within-group effects comparing follow-up against pre-intervention

The overall summary effect on stress at follow-up was moderate (Hedge's $g = -0.611$, 95% CI $[-0.773, -0.449]$, $p = .000$). Figure 5 shows the forest plot of the analysis.

The I^2 statistics for operational stress ($I^2 = 75.132$, $p = .007$), and organisational stress ($I^2 = 82.454$, $p = .003$) were high, and perceived stress ($I^2 = 50.778$, $p = .058$) was moderate. Meta-regression indicated that the moderators of sample size ($Q = 16.88$, $df = 8$, $p = .031$), and research design ($Q = 8.42$, $df = 2$, $p = .016$) were responsible for the

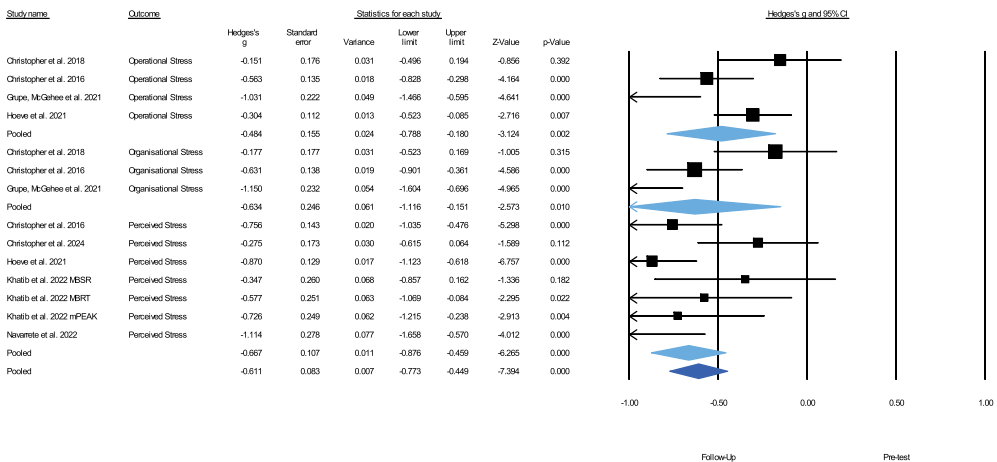


Figure 5. Forest plot of within-group analysis at follow-up on stress.

significant heterogeneity. The meta-regression results are displayed in Table S12 of the Supplementary Materials. Egger's regression test (intercept = -1.606 , $t = .932$, $df = 12$, $p = .370$) indicated publication bias was not detected. Duval and Tweedie's trim-and-fill procedure and inspection of the adjusted funnel plot indicated two missing studies, which altered the effect from Hedge's $g = -.611$, 95% CI $[-.773, -.449]$ to Hedge's $g = -.531$, 95% CI $[-.701, -.361]$, however the direction and magnitude of the effect remained unchanged. Rosenthal's fail-safe N analysis (Rosenthal, 1979) determined that 565 additional studies with null results would be necessary to overturn the observed effect. The funnel plot is presented in Figure S4 of the Supplementary Materials.

Analysis of RQ2

RQ2: Are mindfulness-based interventions effective for reducing mental health outcomes in police officers?

Between-group effects comparing MBI and control groups at post-intervention

As presented in Figure 6, the overall summary effect of MBIs on mental health outcomes at post intervention yielded a small significant effect compared to controls (Hedge's $g = -.273$, 95% CI $[-.427, -.119]$, $p = .001$). The pooled effect sizes of individual outcomes indicated MBIs had a small-moderate effect on anxiety (Hedge's $g = -.398$, 95% CI $[-.725, -.070]$, $p = .017$) and a small effect on sleep disturbance (Hedge's $g = -.236$, 95% CI $[-.458, -.014]$, $p = .037$) in comparison to control groups. While a small effect was indicated for PTSD (Hedge's $g = -.248$, 95% CI $[-.770, .274]$, $p = .351$), depression (Hedge's $g = -.315$, 95% CI $[-.753, .123]$, $p = .159$), and alcohol use (Hedge's $g = -.126$, 95% CI $[-.422, .175]$, $p = .639$) at post-intervention, significance was not reached.

The I^2 value for and depression ($I^2 = 78.445$, $p = .001$) indicated substantial heterogeneity. Heterogeneity was also detected for PTSD ($I^2 = 57.241$, $p = .126$), anxiety ($I^2 = 57.152$, $p = .072$), sleep disturbance ($I^2 = 8.355$, $p = .351$), and alcohol use

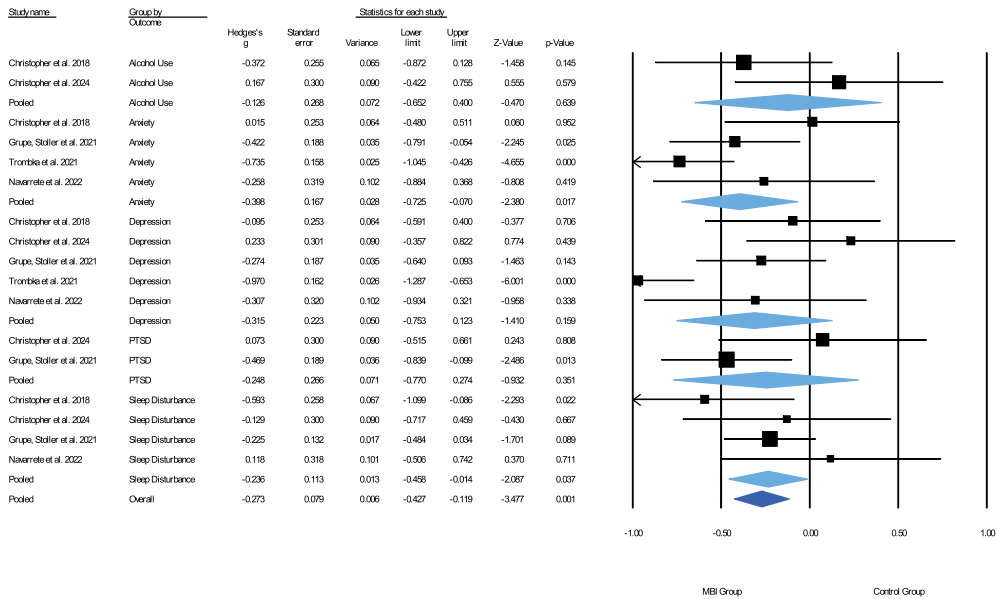


Figure 6. Forest plot of between-group analysis comparing MBI and Control groups at post-intervention on mental health outcomes.

($I^2 = 46.471, p = .172$), but was not significant. Meta-regression indicated the moderators of publication year ($Q = 10.04, df = 3, p = .018$), country ($Q = 24.12, df = 2, p = .000$), sample size ($Q = 29.79, df = 4, p = .000$), and MBI protocol ($Q = 26.76, df = 3, p = .000$) accounted for the heterogeneity. The results of the meta-regression are presented in Table S13. Publication bias was detected by Egger’s regression test (intercept = 2.861, $t = 2.590, df = 15, p = .020$). Duval and Tweedie’s trim-and-fill procedure, along with an examination of the adjusted funnel plot, revealed six potentially missing studies. Inclusion of these studies increased the observed effect size from Hedge’s $g = -.273, 95\% \text{ CI } [-.427, -.119]$ to Hedge’s $g = -.457, 95\% \text{ CI } [-.624, -.202]$. The funnel plot is presented in Figure S5 of the Supplementary Materials.

Between-group effects comparing MBI and control groups at follow-up

As shown in Figure 7, the overall summary effect on mental health outcomes at follow-up was small (Hedge’s $g = -.276, 95\% \text{ CI } [-.434, -.119], p = .001$). The pooled effects show MBIs had a small effect on sleep disturbances (Hedge’s $g = -.388, 95\% \text{ CI } [-.603, -.172], p = .000$) in comparison to control groups at follow-up. Small effects were evident for depression (Hedge’s $g = -.166, 95\% \text{ CI } [-.609, .277], p = .463$), anxiety (Hedge’s $g = -.307, 95\% \text{ CI } [-.671, .058], p = .099$), and alcohol use (Hedge’s $g = .060, 95\% \text{ CI } [-.343, .463], p = .770$), however these effects were not significant.

The I^2 statistics for depression ($I^2 = 77.216, p = .004$) indicated a high degree of heterogeneity. Moderate and low, but not significant heterogeneity was detected for anxiety ($I^2 = 63.219, p = .066$), and alcohol use ($I^2 = 10.809, p = .290$), respectively. A zero I^2 statistic indicated an absence of heterogeneity for sleep disturbances. Meta-regression indicated the moderators of publishing year ($Q = 18.94, df = 2, p < .001$),

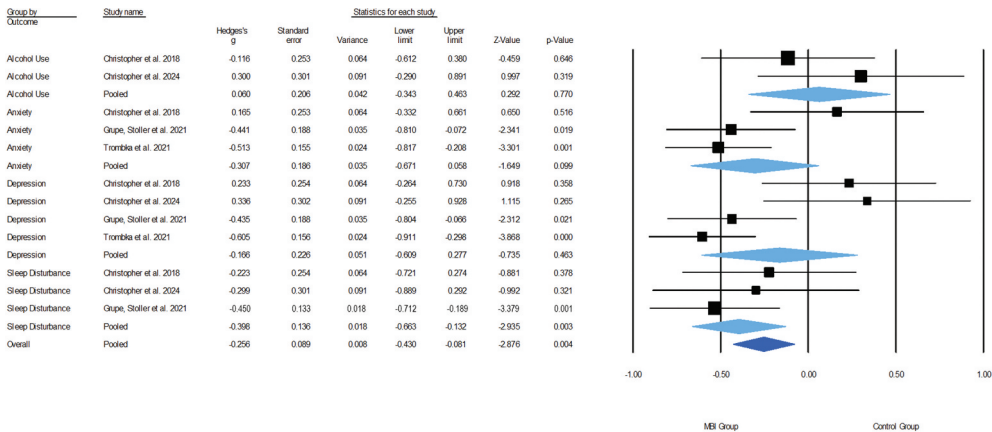


Figure 7. Forest plot of between-group analysis comparing MBI and Control groups at follow-up on mental health outcomes.

sample size ($Q = 19.56$, $df = 3$, $p < .001$), and MBI protocol ($Q = 19.35$, $df = 3$, $p < .001$) were accountable for the significant heterogeneity. The meta-regression results are displayed in Table S14. Egger's regression test indicated publication bias was present (intercept = 4.305, $t = 4.780$, $df = 10$, $p < .001$). Duval and Tweedie's trim-and-fill analysis and inspection of the imputed funnel plot identified two missing studies, which altered the effect from Hedge's $g = -.276$, 95% CI $[-.434, -.119]$ to Hedge's $g = -.309$, 95% CI $[-.497, -.122]$. The funnel plot is displayed in Figure S6 of the Supplementary Materials.

Sensitivity analysis

To determine the degree of robustness of the findings, seven sensitivity analyses were performed. Removing one study with a follow-up period of 6 months (Trombka et al., 2021), and one study which delivered MBI online (Christopher et al., 2024) and did not change the direction or significance of the effect at post-intervention or follow-up on overall mental health. Similarly, removing one study at post-intervention which comprised a NRCT (Navarrete et al., 2022) did not alter the direction or significance of the effect. After removing one study which employed a closely modified version of MBSR (Grupe, Stoller, et al., 2021), the effect of MBIs on overall mental health changed from significant to not significant at post-intervention ($p = .079$), and follow-up ($p = .360$). The results of the sensitivity analyses are shown in Tables S15 – S21.

Within group effects comparing post-intervention against pre-intervention

As shown in Figure 8, the effect of MBIs on mental health outcomes overall was small (Hedge's $g = -.361$, 95% CI $[-.481, -.241]$, $p = .000$). The I^2 statistics for anxiety ($I^2 = 70.636$, $p = .004$), depression ($I^2 = 75.626$, $p = .000$), and sleep disturbance ($I^2 = 58.399$, $p = .047$) indicated heterogeneity was moderate. Heterogeneity was not significant for PTSD ($I^2 = 64.121$, $p = .095$, and alcohol use ($I^2 = 20.313$, $p = .263$). Meta-regression indicated the moderators of sample size ($Q = 27.35$, $df = 8$, $p < 0.001$), and study design ($Q = 4.03$, $df = 1$, $p = .046$) accounted for the heterogeneity. Publication bias

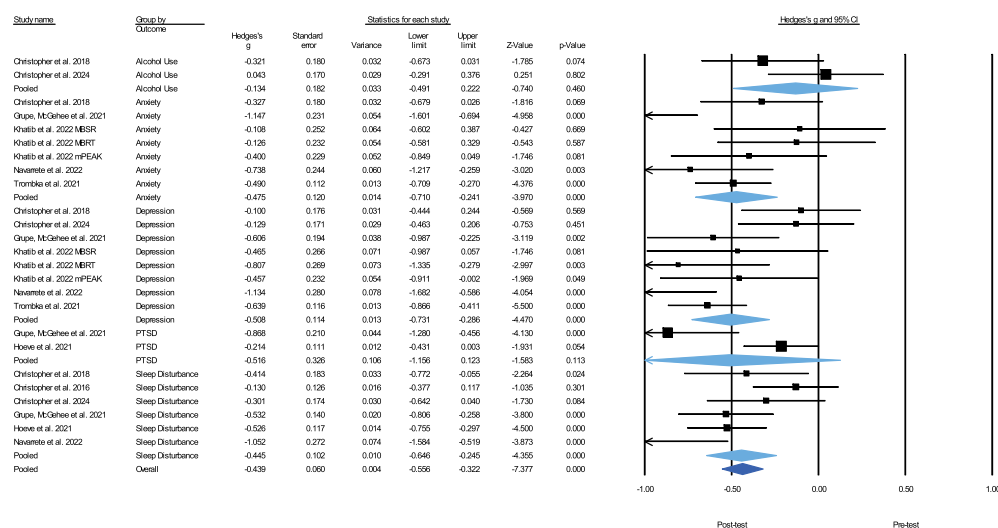


Figure 8. Forest plot of within-group analysis at post-intervention on mental health outcomes.

was not identified (intercept = -2.366 , $t = 1.957$, $df = 20$, $p = .064$). The results of the meta-regression are presented in Table S22. Duval and Tweedie's trim-and-fill procedure, along with an examination of the adjusted funnel plot, revealed two potentially missing studies. Inclusion of these studies increased the effect size from Hedge's $g = -.361$, 95% CI $[-.481, -.241]$ to an effect size of Hedge's $g = -.405$ (95% CI $[-.527, -.284]$). Figure S7 of the Supplementary Materials displays the funnel plot.

Within-group effects comparing follow-up against pre-intervention

As presented in Figure 9, the overall summary effect showed a small effect of MBI on mental health outcomes (Hedge's $g = -.361$, 95% CI $[-.481, -.241]$, $p = .000$). The I^2 statistics for anxiety ($I^2 = 70.636$, $p = .004$), and depression ($I^2 = 75.626$, $p = .003$) suggested a high degree of heterogeneity, and a moderate degree was detected for sleep disturbances ($I^2 = 58.399$, $p = .047$). Moderate and low but not significant heterogeneity was detected for PTSD ($I^2 = 64.121$, $p = .095$), and alcohol use ($I^2 = 20.313$, $p = .263$), respectively. Meta-regression indicated the moderators of sample size ($Q = 27.35$, $df = 8$, $p < .001$), and research design ($Q = 4.03$, $df = 1$, $p = .045$) accounted for the significant heterogeneity. Table S23 presents the results of the meta-regression. As indicated by Egger's regression test (intercept = -2.366 , $t = 1.957$, $df = 20$, $p = .064$), publication bias was not detected. Duval and Tweedie's trim-and-fill analysis and inspection of the funnel plot identified three missing studies, which altered the effect from Hedge's $g = -.361$, 95% CI $[-.481, -.241]$ to Hedge's $g = -.335$, 95% CI $[-.478, -.192]$, however the magnitude and direction of the effect did not change. The fail-safe N values indicated 653 additional studies would be required to nullify the effect. The funnel plot is presented in Figure S8 of the Supplementary Materials.

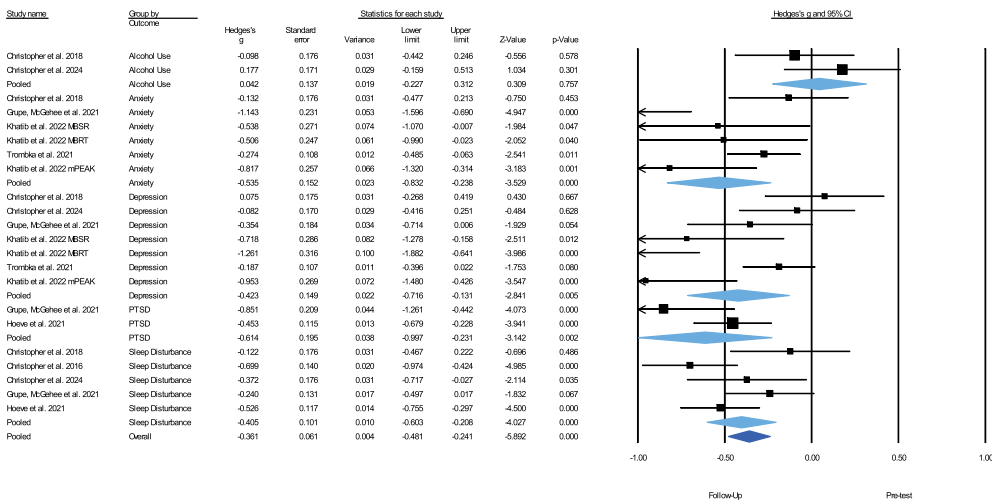


Figure 9. Forest plot of within-group analysis at follow-up on mental health outcomes.

Discussion

The aim of this systematic review and meta-analysis was to evaluate whether MBIs are effective for reducing stress and improving mental health outcomes in police officers. Eleven studies with 967 participants were included.

Results from the between- and within-group analyses provide evidence towards MBIs being effective for reducing stress and improving mental health in police officers. Between-group comparisons indicated a small effect of MBIs favouring the MBI group compared to controls on overall stress and mental health outcomes at post-intervention, however, only a significant effect on mental health remained at follow-up. The within-group comparisons show that MBIs demonstrated small to moderate effects for reducing overall stress and improving general mental health in police officers at post-intervention. These effects increased to moderate at follow-up. These findings are generally consistent with other meta-analyses which have found MBIs to be effective for reducing stress and improving mental health outcomes in other populations, however the effects of MBIs in other populations tend to be larger than the observed effects in this study (Li, Nannestad, et al., 2024; Ong et al., 2024; Qin et al., 2024; Xunlin et al., 2020).

Between-group comparisons detected small to moderate effects of MBIs on operational, organisational and perceived stress, and cortisol awakening response at post-intervention and follow-up, however the effects did not reach significance. The analysis showed an aggregated significant reduction in overall stress at post-intervention. This result underscores the enhanced statistical power afforded by the meta-analytic approach, which enables the detection of cumulative effects that might remain undetected in smaller-scale studies and individual outcomes (Borenstein et al., 2021). Two of the included studies (Christopher et al., 2024, 2018) were feasibility trials characterised by relatively small sample sizes. The authors of both studies suggested that the non-significant results at post-intervention might be attributed to the limited sample sizes, which in turn reduced statistical power. The sample included in our meta-analysis was also

small, therefore significance may not have been reached despite small to moderate effects being present. The absence of significant effects on individual stress outcomes could also be due to the persistent and chronic nature of stressors in police work. Other meta-analytic reviews evaluating MBIs in other comparable high-stress populations (Ong et al., 2024; Rogerson et al., 2024), have also found small and/or non-significant effects on stress. Addressing high-baseline stress levels among police, as well as the intensity and complexity of stressors, may require more time, and sustained daily mindfulness practice. Moreover, the interventions employed across the included RCTs were modified for use in police, with practices focusing on police work. These modifications may have restricted participant's practice of mindfulness to the policing environment, subsequently impeding the therapeutic benefit of the interventions. Regular and sustained daily mindfulness practice forms a core component of mindfulness training (Kabat-Zinn, 2013; Parsons et al., 2017), therefore the opportunity for participants to practice mindfulness consistently beyond the policing environment may have inhibited the effectiveness. Furthermore, qualitative data from Fitzhugh et al. (2024) study suggests that police officers may be hesitant to practice mindfulness in the workplace due to embarrassment and stigma, therefore teaching mindfulness practices which can be implemented and practiced beyond the policing environment, such as MBSR (Kabat-Zinn, 1982), may provide a more effective treatment for stress in police. Consistent with this, the lack of significant overall effect on stress observed at follow-up may be due to low rates of participant engagement in mindfulness practices beyond the intervention. The included studies reported low rates of engagement in mindfulness by participants during the follow-up period. As such the longer-term therapeutic value of the programs may have been impacted due to low ongoing practice rates. The extent of formal and informal mindfulness practice has been shown to be positively associated with intervention outcomes (Parsons et al., 2017), however other studies have found no association between engagement in daily practice and changes in stress (Grupe, Stoller, et al., 2021).

Our meta-analysis also evaluated the effect of MBIs on changes in biomarkers of stress in police, specifically changes in the cortisol awakening response (CAR). While only a small trend-level reduction in CAR was observed favouring the MBI treatment groups at post-intervention (Hedges $g = -.289$, $p = .055$), this may indicate normalisation of the stress-hormone, cortisol. The results are similar in magnitude to the findings of Rogerson et al. (2024) meta-analysis examining the effectiveness of stress management interventions on cortisol awakening response levels in other high-stress populations. Only one RCT (Grupe, Stoller, et al., 2021) provided follow-up data for CAR, therefore we were unable to perform a meta-analysis for changes in CAR beyond post-intervention for between-group comparisons. The results of Grupe, Stoller, et al. (2021) study however indicated participants in the MBI intervention group had significantly lower CAR at follow-up (3 months) relative to waitlist participants. These findings suggest the benefits of MBIs may emerge more slowly over time, and may buffer against the consequences of chronic stress. Other biological markers of stress have also been studied. One study examined the effect of MBIs on hair cortisol concentrations (HCC; Grupe, Stoller, et al., 2021), however no significant effect was observed at post-intervention or 3-month follow-up. This result may be attributed to divergence of cortisol levels across hair samples, where high variations in HCC exist dependent on cumulative traumatic load and stress, PTSD, comorbid mental health issues, and childhood adversity (Koumantarou Malisiova et al., 2021). One RCT (Krick & Felfe,

2020) assessed the effect of MBIs on heart rate variability (HRV) with results indicating a positive effect of MBIs on HRV at post-intervention. The combined findings may underscore the complex interactions between chronic stress psychopathology, as well as the potential of MBIs to alter biological markers of stress in police. Given the link between pathological functioning of biological stress systems and adverse psychological health outcomes (Creswell & Lindsay, 2014; Finlay et al., 2022; McEwen, 2017), further research assessing the effect of MBIs on biomarkers of stress among policing samples is needed to comprehensively evaluate the impact of MBIs on biological stress indicators. Furthermore, the effects of MBIs on biological stress markers may emerge over time (Grasmann et al., 2023), therefore, future studies involving police should incorporate medium and long-term follow-up measures of stress biomarkers to assess sustained or delayed MBI treatment impacts.

Relative to controls, MBI groups demonstrated small-moderate reductions in anxiety at post-intervention. A number of studies suggest that MBIs, particularly MBIs employing focused meditation practices, recruit multiple brain regions involved in attentional control (Ganesan et al., 2022; Tang et al., 2015) which are associated with decreases in anxiety, and threat responses (Creswell & Lindsay, 2014; Eysenck et al., 2007; Williams et al., 2024). Small reductions in sleep disturbances were also evident at post-intervention, with effects strengthening at follow-up. Despite the significant improvements in anxiety and sleep disturbances, our meta-analysis found no significant differences in depression between MBI treatment groups and controls at both post-intervention and follow-up, which contrasts with the moderate effects reported by Withrow et al. (2024). This discrepancy is likely attributable to Withrow et al. (2024) incorporating outcomes from Krick and Felfe's (2020) study, which did not specifically measure depression or anxiety. Instead, Krick and Felfe (2020) assessed negative affect using the Positive and Negative Affect Schedule-X (PANAS-X; R ocke & Gr uhn, 2003), and the Psychological Strain Scale (Irritation Scale; Mohr et al., 2005) which also include emotions such as fear, anxiety, and hostility, therefore the effects observed may be attributable to changes in related symptoms rather than depression and anxiety specifically.

The results of our study suggested a large effect of MBIs on PTSD symptoms at post-intervention, however the effect did not reach significance. Only one study reported data for this could indicate that MBIs may be effective for reducing emotional reactivity, and physiological arousal associated with PTSD in police. Moreover, mindfulness training may serve to reduce the full range of negative cognitions, moods, trauma-related reactivity, sleep disturbances and anxiety to support successful engagement in other PTSD-specific interventions (Nagy et al., 2020). Only one study provided data for PTSD at follow-up (Grupe, Stoller, et al., 2021), therefore a meta-analysis at follow-up could not be performed.

The cumulative findings of our meta-analysis suggest that while some stressors and emotional experiences may not change, the effects on behaviours, such as sleep disturbances, may shift. This shift in reaction to stress or emotions is aligned with mindfulness-based practices, and has been seen in other studies examining mindfulness and MBIs (Nagy et al., 2020). To this end, while stressors may remain, mindfulness may have increased, thereby moderating and mediating the impact of stress on the wellbeing of officers. Future studies should examine the mechanisms of mindfulness among officers, to provide a more comprehensive understanding the benefits of MBIs for police.

Limitations and future research

There are some limitations that exist in both the included studies and our meta-analysis. Firstly, only a small number of studies were eligible for inclusion for between-group analyses, therefore the ability to detect a significant effect may be reduced. Future RCTs should attempt to increase sample sizes to enable significant effects to be detected. Secondly, the inclusion of NRCTs in our between-group analyses may introduce potential biases or confounding factors that could not be fully controlled. However, the robustness of our findings, as demonstrated by the sensitivity analysis, supports the validity and reliability of our conclusions. Thirdly, limited studies incorporated biological measures of stress as outcome variables, and few studies provided follow-up data. Given the association between the persistent activation of the stress response and adverse psychological health outcomes (Finlay et al., 2022; McEwen, 2017), measurement of biological markers, particularly those implicated in chronic stress, is recommended. Moreover, other studies, including meta-analytic reviews in other populations (Grasmann et al., 2023; Rogerson et al., 2024), have found positive outcomes or greater effects of MBIs at longer-term follow-up, future studies should incorporate longer-term follow-up measures into the design to determine the long-term benefits of MBIs for police. Researchers implementing MBIs within policing environments should also consider including booster sessions to determine whether ongoing mindfulness training sessions improve medium- and long-term treatment effects on stress. Future studies assessing the effects of MBIs on stress and mental health outcomes, including sleep over extended time periods are therefore warranted. Fourth, the results of the meta-regression analyses demonstrated substantial levels of heterogeneity across some outcomes, implying that the effects of MBIs on stress and mental health outcomes are higher in some police populations and lower in others (Borenstein, 2019). Caution should be utilised when generalising the findings from this meta-analysis to all police populations. Fifth, some evidence of publication bias was observed for certain outcomes, which might influence the reliability of the results and should be considered when interpreting the findings. Sixth, most interventions employed within the primary studies were adapted to more closely align with policing culture, which may be counterintuitive. Results of the sensitivity analyses indicate that after removing the study which employed a closely modified MBSR protocol (Grupe, Stoller, et al., 2021), the significant effects on overall mental health disappeared at post-intervention and follow-up, suggesting interventions more closely aligned to MBSR or protocols which do not focus on the policing environment and stressors specifically, may be more effective for improving mental health in police officers. Future research should seek to evaluate the effectiveness of the original MBSR curriculum (Kabat-Zinn, 1982, 2015) on stress in police, which can be more broadly applied to daily life. Lastly, there is a significant lack of reporting on the adverse effects of MBSR in the individual studies, particularly concerning adverse psychological and social impacts resulting from the intervention. This absence of data makes it challenging to fully understand the potential risks associated with the intervention (Li, McIntyre, et al., 2024). Future research should prioritise systematic investigations of potential adverse effects and factors contributing to adverse impacts, including experiences of trauma in childhood, to offer a more thorough understanding of the

intervention's safety profile for use in police settings.

Recommendations for police agencies

Police agencies are encouraged to establish research partnerships with academic institutions, mental health professionals, and other organisations to deepen the understanding of MBIs within the unique context of policing. These collaborations should focus on leveraging academic expertise including the specialised knowledge and methodological skills of researchers to design and conduct high-quality studies, as well as securing funding through grant applications and sharing resources to support rigorous studies exploring the effectiveness of MBIs in enhancing officer wellbeing. Focus should also be given to assessing the safety profile of MBIs for use with police officers to ensure interventions are both effective and safe with police. Furthermore, collaborative efforts should aim to implement pilot programs to test the feasibility and effectiveness of MBIs across various countries, jurisdictions, and sub-populations of police (specialised units, recruits, high versus low stress and mental health symptomology) to understand the benefits of MBIs among policing populations, and established an understanding of who, under what conditions, are MBIs effective and safe. It is also recommended that police agencies evaluate existing mindfulness programs and initiatives which have been implemented by organisations through empirical and systematic assessment to support refinement, improvement of programs, and to establish the safety profile of existing initiatives and interventions. By fostering research collaborations, agencies can contribute to the broader evidence-base for the use of MBIs among police and ensure that resources are directed toward programs with proven efficacy and safety.

Conclusion

Our study offers important implications for future research, practice and policy. The results of the current review and meta-analysis show reductions in overall stress and improved overall mental health outcomes among police following MBI interventions. Findings from the study provide supportive evidence that MBIs may have a positive cumulative impact on officer wellbeing and can be utilised to support police officers manage stress and enhance their mental health while serving in the profession. Our meta-analysis provides preliminary support for police agencies to establish partnerships with research and academic institutions to rigorously examine the use of MBIs, both independently and alongside existing mental health support frameworks, and to guide leadership decision-making processes to enhance occupational health policy to improve officer wellbeing.

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Notes on contributors

Jennifer Gaskin is a PhD candidate within the College of Healthcare Sciences at James Cook University, Australia. Her research focuses on stress, mental health, mindfulness, and policing. She obtained a Bachelor of Criminology and Bachelor of Psychological Science at Deakin University, Australia, and Advanced Post-Graduate Diploma of Psychological Practice at Charles Sturt University, Australia. She has worked extensively across the Australian criminal justice system, including police, courts, and corrective services.

Wendy Li is an experienced researcher in the field of mental health and wellbeing employing mindfulness interventions and has published extensively in the areas of mental health, and mindfulness. She is a qualified Mindfulness-Based Stress Reduction (MBSR) teacher certified by Brown University, and founder and leader of the Mindfulness Lab at James Cook University. Wendy completed her PhD at the University of Waikato, New Zealand. She is currently an Associate Professor in the Department of Psychology at James Cook University, Australia.

Liza van Eijk's background is in neuropsychology and neuroimaging. Her research focuses on improving early detection, diagnosis as well as treatment outcomes for several conditions including stress and mental health disorders, by utilising early markers in prediction models. She obtained her PhD at the University of Queensland, Australia, and Research Master in Cognitive Neuropsychology at the VU University, The Netherlands, cum laude, which included a research internship at Cambridge University and a clinical neuropsychology internship. She is a Senior Lecturer in the Department of Psychology at James Cook University, Australia.

Zoltan Sarnyai is a medically trained PhD neuroscientist with internationally recognised expertise in stress and mental health research. He was trained at Harvard Medical School and Rockefeller University in the USA and was on the faculty of the University of Cambridge, UK. He is currently Professor within the College of Medicine and Dentistry at James Cook University, Australia.

ORCID

Jennifer Gaskin  <http://orcid.org/0009-0008-5985-527X>

Wendy Wen Li  <http://orcid.org/0000-0002-0056-6130>

Liza van Eijk  <http://orcid.org/0000-0001-8254-1019>

Zoltan Sarnyai  <http://orcid.org/0000-0002-2380-7902>

Authorship contribution statement

Jennifer Gaskin significantly contributed to the conceptualisation, formal analysis, investigation, methodology, project administration, visualization, writing – original draft, and writing (review and editing) of this paper. Wendy Li contributed to the conceptualisation, formal analysis, investigation, methodology, writing (review and editing) of this paper, and provided supervision to Ms Gaskin. Both Liza van Eijk and Zoltan Sarnyai contributed to the conceptualisation, methodology, writing (review and editing) of this paper, and provided supervision to Ms Gaskin.

Data availability statement

All data generated or analysed during this study are included in this published article and its supplementary material files.

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